

## Supplementary material of “Dating multiple change points in the correlation matrix”

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### 1 Proofs

#### *Proof of Theorem 1*

The proof is similar to the proof of Theorem 1 in Galeano and Wied (2014). Denote  $P_{l_1, l_2}^{**}(z) := P_{l_1, l_2}^{**, i, j}(z)_{1 \leq i < j \leq p}$ . Consider a fixed pair  $(i, j)$ ,  $1 \leq i < j \leq p$  and the process

$$\begin{aligned} B_{\eta(l_1), \tau(l_2)}^{i, j}(z) &:= \frac{\tau(z) - \eta(l_1) + 1}{\tau(l_2) - \eta(l_1) + 1} P_{\tau(z), \eta(l_1), \tau(l_2)}^{i, j} \\ &:= \frac{\tau(z) - \eta(l_1) + 1}{\tau(l_2) - \eta(l_1) + 1} \left( \hat{\rho}_{\eta(l_1), \tau(z)}^{i, j} - \rho_{i, j} \right) - \frac{\tau(z) - \eta(l_1) + 1}{\tau(l_2) - \eta(l_1) + 1} \left( \hat{\rho}_{\eta(l_1), \tau(l_2)}^{i, j} - \rho_{i, j} \right) \end{aligned}$$

with  $\rho_{i, j} := \frac{(m_{XX'})_{i, j} - \mu_i \mu_j}{\sqrt{\sigma_i^2 \sigma_j^2}}$ . Then, by arguments based on the strong law of large numbers, one can show that, uniformly for  $z \in [l_1 + \varepsilon, l_2]$  (for a small  $0 < \varepsilon < l_2 - l_1$ )  $B_{\eta(l_1), \tau(l_2)}^{i, j}(z) \Rightarrow_d P_{l_1, l_2}^{**, i, j}(z)$ . By means of arguments based on Theorem 4.2 in Billingsley (1968) one gets convergence uniformly in  $[l_1, l_2]$ . Then, uniformly in  $[l_1, l_2]$ ,

$$B_{\eta(l_1), \tau(l_2)}(z) \rightarrow_{a.s.} P_{l_1, l_2}(z).$$

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The theorem follows by applying the argmax continuous mapping theorem (Kim and Pollard, 1990, Theorem 2.7).  $\blacksquare$

*Proof of Theorem 2*

We assume without loss of generality that there is only one change point in  $k_0 = [Tz_0]$  and that  $[l_1, l_2] = [0, 1]$ . Denote  $P_{0,1}(z) =: P(z)$ . Then  $P(z)$  has a unique maximum in  $z_0$ . Similarly as in the proof of Proposition 2 in Bai (1997), we show that

$$P_{M,T} := P \left( \max_{|k-k_0| > M} B_{1,T} \left( \frac{k}{T} \right) - B_{1,T} \left( \frac{k_0}{T} \right) \geq 0 \right)$$

becomes small for large  $M$  and  $T$ . That means that, for every  $\varepsilon > 0$ , there is a  $M > 0$  and a  $T_0 > 0$  such that, for all  $T > T_0$ ,  $P_{M,T} < \varepsilon$ .

Now,  $B_{1,T} \left( \frac{k}{T} \right) - B_{1,T} \left( \frac{k_0}{T} \right) \geq 0$  is equivalent to

$$\left( B_{1,T} \left( \frac{k}{T} \right) - P \left( \frac{k}{T} \right) \right) - \left( B_{1,T} \left( \frac{k_0}{T} \right) - P \left( \frac{k_0}{T} \right) \right) + \left( P \left( \frac{k}{T} \right) - P \left( \frac{k_0}{T} \right) \right) \geq 0.$$

Assume for the moment  $k > k_0$  and that the standard deviations of all random variables are equal to 1. (Divide each component of  $P(\cdot)$  by the standard deviations if the latter assumption is not fulfilled.) We multiply the whole equation with  $T/(k - k_0)$  and denote

$$A_1(k, k_0, T) = \frac{T}{k - k_0} \left( \left( B_{1,T} \left( \frac{k}{T} \right) - P \left( \frac{k}{T} \right) \right) - \left( B_{1,T} \left( \frac{k_0}{T} \right) - P \left( \frac{k_0}{T} \right) \right) \right)$$

$$A_2(k, k_0, T) = \frac{T}{k - k_0} \left( P \left( \frac{k}{T} \right) - P \left( \frac{k_0}{T} \right) \right).$$

Now, we use several observations in order to argue that the asymptotic behavior of  $A_1(k, k_0, T)$  can be reduced to the behavior of

$$\sum_{1 \leq i < j \leq p} \frac{1}{k - k_0} \sum_{t=k_0}^k (X_{t,i} X_{t,j} - \mathbb{E}(X_{t,i} X_{t,j})).$$

This quantity is then arbitrarily small by the law of large numbers for sufficiently large  $M$ . The observations are the following:

1.  $A_1(k, k_0, T)$  can be regarded as the sum of  $p(p-1)/2$  components and each component can be treated separately.
2. For large  $T$ , with high probability and uniformly in  $z \in \{k_0/T, k/T\}$ , all components of  $TP_{\tau(z),1,T}$  and  $TP_{0,1}^*(z)$  have the same sign so that we consider the values without having applied the absolute value function.
3. For large  $T$ , with high probability and uniformly in  $z \in \{k_0/T, k/T\}$ , the successive variances in the denominators of the components of  $A_1(k, k_0, T)$  are equal to their theoretical counterparts which are the same for  $z = k/T$  and  $z = k_0/T$ , respectively.

4. In the numerators of the components of  $A_1(k, k_0, T)$ , we have expressions like

$$\frac{1}{k - k_0} \left( \left( \sum_{t=k_0}^k X_{t,i} X_{t,j} \right) - T \left( P_{i,j} \left( \frac{k}{T} \right) - P_{i,j} \left( \frac{k_0}{T} \right) \right) - \left( \frac{1}{k} \sum_{t=1}^k X_{t,i} \sum_{t=1}^k X_{t,j} - \frac{1}{k_0} \sum_{t=1}^{k_0} X_{t,i} \sum_{t=1}^{k_0} X_{t,j} - (k - k_0) \frac{1}{T} \sum_{t=1}^T X_{t,i} \frac{1}{T} \sum_{t=1}^T X_{t,j} \right) \right).$$

Here,  $P_{i,j}(\cdot)$  are the components of  $P(\cdot)$ .

5. It holds

$$T \left( P_{i,j} \left( \frac{k}{T} \right) - P_{i,j} \left( \frac{k_0}{T} \right) \right) = (k - k_0) \mathbb{E}(X_{k,i} X_{k,j}).$$

Then, after some tedious calculations, one sees that  $A_1(k, k_0, T)$  is a random variable such that, for all  $\varepsilon > 0$  and all  $\eta > 0$ , there is a  $M > 0$  such that  $\mathbb{P}(|A_1(k, k_0, T)| > \varepsilon) < \eta$  for  $k > k_0 + M$  and  $T > T_0$ . This means that  $A_1(k, k_0, T)$  is arbitrarily small whenever  $T$  and  $M$  are large. On the other hand,  $A_2(k, k_0, T)$  does not converge to zero:  $P\left(\frac{k}{T}\right) - P\left(\frac{k_0}{T}\right)$  is a finite sum of linear functions in  $k$  with negative slope (see Figure 1 in Galeano and Wied (2014)) so that it is equal to  $C\left(\frac{k}{T} - \frac{k_0}{T}\right)$  for a  $C < 0$  by Taylor's formula. Multiplied with  $T/(k - k_0)$ , the expression is equal to  $C$ . Then, with large probability,  $A_1(k, k_0, T) + A_2(k, k_0, T)$  is strictly negative. For  $k < k_0$ , the argument is similar and the theorem is proven. ■

*Proof of Theorem 3*

Denote  $Q_T^{l_1, l_2} := \sup_{z \in [l_1, l_2]} A_{\eta(l_1), \tau(l_2)}(z)$  the test statistic calculated from data from  $\eta(l_1)$  to  $\tau(l_2)$ . Moreover, let  $B \in \mathbb{N}$  a fixed number of bootstrap repetitions. Now, by Theorem 1, it holds

$$\frac{1}{a_T^k} B_{\eta(l_1), \tau(l_2)}(z) \rightarrow_p \infty$$

for any sequence  $a_T^k = o(\sqrt{T})$  if there is a change point in the interval  $[l_1, l_2]$ . Moreover, due to Assumption 9.b, the eigenvalues of  $\hat{E}_{\eta(l_1), \tau(l_2)}$  are stochastically bounded. Consequently, the eigenvalues of  $\hat{E}_{\eta(l_1), \tau(l_2)}^{-1/2}$  (remember that we assume its existence) are bounded away from zero and the matrix is positive definite. Therefore,

$$\frac{1}{a_T^k} Q_T^{l_1, l_2} \rightarrow_p \infty$$

(with respect to the measure  $\mathbb{P}^\times$ ) for any sequence  $a_T^k = o(\sqrt{T})$  if there is a change point in the interval  $[l_1, l_2]$ . By Theorem 2, we moreover have

$$\frac{1}{a_T^k} Q_T^{\hat{z}_i, \hat{z}_{i+1}} \rightarrow_p \infty \quad (1)$$

(with respect to the measure  $P^\times$ ), where  $\hat{z}_i$  and  $\hat{z}_{i+1}$  for  $i \in \mathbb{N}_0$  are two estimated change points in one of the iterations of the algorithms, as long as there is a change point in the interval  $[z_i, z_{i+1}]$ . This follows from the fact, that, by Theorem 2,

$$Q_T^{\hat{z}_i, \hat{z}_{i+1}} - Q_T^{z_i, z_{i+1}} = o_{P^\times} \left( \frac{1}{\sqrt{T}} \right).$$

Moreover, with the same argument and with the results under the null hypothesis from Wied (2015),

$$Q_T^{\hat{z}_i, \hat{z}_{i+1}} = O_{P^\times}(1) \quad (2)$$

if there is no change point in the interval  $[z_i, z_{i+1}]$ .

Then, due to (1), it holds  $\lim_{T \rightarrow \infty} P^\times(\hat{\ell} < \ell) \rightarrow 0$  and due to (2), it holds  $\lim_{T \rightarrow \infty} P^\times(\hat{\ell} > \ell) \rightarrow 0$ , which proves the theorem (compare the proof of Proposition 11 in Bai, 1997). ■

#### *Proof of Theorem 4*

In this proof, we combine ideas of the proof of Theorem 3 in Galeano and Wied (2014) and of Theorem 3 in Wied (2015). Note that  $\hat{z}^* = \operatorname{argmax}_{l_1 \leq z \leq l_2} B_{\eta(l_1), \tau(l_2)}^*(z)$  with

$$B_{\eta(l_1), \tau(l_2)}(z) := \frac{\tau(z) - \eta(l_1) + 1}{\sqrt{\tau(l_2) - \eta(l_1) + 1}} \left\| P_{\tau(z), \eta(l_1), \tau(l_2)} \right\|_1.$$

Then, one can show by means of the extended functional delta method from Wied et al (2012) that the process  $\frac{\tau(z) - \eta(l_1) + 1}{\sqrt{\tau(l_2) - \eta(l_1) + 1}} P_{\tau(z), \eta(l_1), \tau(l_2)}$  converges to the process

$$E^{1/2} W^{\frac{p(p-1)}{2}}(z) - E^{1/2} W^{\frac{p(p-1)}{2}}(l_1) - \frac{z - l_1}{l_2 - l_1} E^{1/2} \left( W^{\frac{p(p-1)}{2}}(l_2) - W^{\frac{p(p-1)}{2}}(l_1) \right) + P_{l_1, l_2}^{**}(z).$$

This limit process has a unique maximum  $P$ -almost surely (Kim and Pollard, 1990, Lemma 2.6) and then the theorem follows by the  $\operatorname{argmax}$  continuous mapping theorem (Kim and Pollard, 1990, Theorem 2.7). ■

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